# Fertility, Reproductive Health and Reproductive History 

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#### Abstract

In one of the first studies ever on the subject we have reported a higher lifetime mortality rate for subfertile and sterile men without co-morbidity over all age groups in comparison to fertile men. The objective of this retrospective cohort study is to explore the association between male fertility and life time mortality. Little is known from literature reviews and own research about eventual biological mechanisms behind these mortality differentials. Here we go on reporting on a survey of surviving subjects and proxys on life time morbidity and reproductive biography details which may give additional information. Databases are general and semen parameters of 1375 patients attending the andrological service at Marburg University Hospital in 1949 and later and data from a core interview with 944 survivors and a proxy interview with 431 surviving family members. Results may indicate a protective effect of cohabitating with children (own/foster/adopted/step) for subfertile men.

\section*{Introduction}

Male infertility is a frequent problem with a complex aetiology. In many industrialized countries about $15 \%$ of all couples trying to become pregnant are infertile giving a waiting time of 12 month.' Although the research on reproductive health gains in importance, one third of the aetiology of fertility disorders remains unclear. Particularly with regard to the hypothesis of declining sperm quality in Western countriesii the research on determinants of reproductive failures is receiving increasing attention. Thus there are a lot of studies which investigate the relationship between risk factors and male infertility. In most of them infertility is referred to as a multifactorial disease and the research interest is on investigating risk factors causing infertility.

Whereas most studies are concerned with the causes of male infertility, there are only a few studies which investigate the outcomes of male infertility on morbidity and/or mortality. Based on the results of a previous published study ${ }^{\text {iii }}$ in which surprisingly a higher mortality risk was found for subfertile men without a specific co-morbidity or previous disease in comparison to fertile men over all age groups (see Figure 1), the target of the current study is to investigate this association between fertility and morbidity on an extended database by linking information from medical records on semen parameters with interview data to provide additional and detailed information for explaining that association.




Figure 1: Survival functions of fertile and subfertile men; early natal-cohorts (In: Groos S: Lebenszeit-Mortalität von Männern mit normalen und subnormalen Spermienkonzentrationen. Diss., Marburg 2005 (p. 60))

## Conceptional Framework

## Objectives

As a plausible biological mechanism for the higher mortality and shorter life span could not be found by literature review yet, we consider the following possibilities:
$>$ a specific comorbidity or specific noxious agents with unfavourable influences on the spermatogenesis were already existing at the time of medical examination for a number of cases with subnormal sperm findings but were not diagnosed or documented (confounder);
> men with subnormal sperm findings had a different life course: because of the higher prevalence of childlessness more unstable partnerships, more frequent changes of employment and residence and a more risky behavior (drift);
$>$ intact fertility is per se a life prolonging specific disposition (direct specific causation);
> intact fertility is - because of a more frequent successful reproduction - a life prolonging specific disposition (indirect specific causation).

## Hypotheses

To explain which of the possibilities come into question, we do not only consider the cumulative mortality in dependence of the sperm concentration but also
$>$ the life time morbidity of the survivor and the deceased cases and
> the reproductive biography of these cases: low sperm density does not exclude reproduction as well as intact fertility is only one of many conditions of actual reproduction.

For this reason we want to test the following hypotheses:
$>$ The morbidity profile of men with oligozoospermia and azoospermia is different from that of men with normozoospermia.
$>$ The prevalence of morbidity of men with oligozoospermia and azoospermia is different from that of men with normozoospermia. The latter will be rarely and less (critically) ill.
> Fertility and fecundity interact with morbidity: by the same fertility status childless men will have a higher and/or a different morbidity as men with children.
> The life course trajectory has an impact on morbidity/mortality. Stresses and strains in marriage/partnership in consequence of fertility disorders are the real cause.


Figure 2: Causal model about the association between fertility and morbidity

In our study we conceptualize fertility not as a typical andrological disease for which risk factors should be identified but rather as a risk factor or exposure variable for male life time morbidity. In this context we refer to the term exposure as a contact of an individual with an agent through any medium or environment. ${ }^{\text {Iiv }}$ Applying exposure to external and internal agents, fertility status can be understood as both - a biologic agent in the body (disorders of the spermatogenesis) as well as a social agent - an incriminatory event in the live time.

Thus, by investigating the association between fertility and life time morbidity direct and indirect effects of involuntary childlessness on male morbidity should be distinguished. In the first case infertility can be considered as a disease per se with objective physical symptoms (abnormal semen characteristics) which manifest in higher morbidity resulting in higher mortality risk. As an example we can take studies on the association between semen characteristics and testicular cancer in which a higher risk of testicular cancer could be found for men with abnormal semen characteristics. ${ }^{\vee}$ In the
second case infertility can be referred to as a stressful life event for the men involved and thus have negative indirect implications for the morbidity by influencing other areas of life associated with ill health. In this context the investigation of the life course of patients after infertility diagnosis is of great importance. Aspects like reproductive biography (type of infertility - primary and secondarypaternity and their alternative forms), partnership history (marital status) and health related behaviour should be taken into consideration by exploring the relationship between fertility status and morbidity.

## Data and Research Methods

## Data

Database are medical records of 2305 men in couples with fertility problems who had a semen analysis done at the fertility and sterility office of the department of andrology at the University Hospital in Marburg during 1949 and 1985 and which were born before 1938. After applying excluding criteria like being born after 31st December 1937, having foreign citizenship, missing values in sperm concentration and pre-existing diseases to assure that at the time of medical examination impairment of spermatogenesis was not the consequence of an existing disease, we have a sample size of 1375 cases, for which the vital status through the registration office could be found out. Thereof 944 cases are still alive and 431 cases are already deceased. To provide additional information explaining the association between sperm concentration and lifetime morbidity a core interview with 944 living cases and an exit interview with 431 surviving family members of deceased cases is conducted on the basis of pretested standardized questionnaire to survey information on general health status, history of diseases, family and partnership history, health related behaviour etc.

## Methods

1. The Estimation of Odds Ratios

In a first step we want to estimate the Odds for fertile and subfertile men according their risks in morbidity. For locating the overestimation of the truth OR therefore two different methods are used for approximating the KI.

Approximative KI of Woolf (1955)

$$
\begin{equation*}
\mathrm{KI}(\mathrm{OR})=\left[\mathrm{OR} \text { est } \cdot \exp \left\{ \pm \mathrm{u}_{1-\alpha / 2} \cdot \sqrt{\frac{1}{\mathrm{n}_{11}}+\frac{1}{\mathrm{n}_{10}}+\frac{1}{\mathrm{n}_{01}}+\frac{1}{n_{00}}}\right\}\right] \tag{1}
\end{equation*}
$$

if estimation is incumbent on Gaussian distribution.
In using the difference of two Logits this interval is also called as Logit-Limits. The most disadvantage of Woolf's approximation is that in cases with small cell frequencies it can't be excepted that the estimation for KI leads to biases results.

Alternative in context to small epidemiological studies it's possible to use the asymptotic KI of Miettinen (1976). This method is a combination of the approximation of $X^{2}$ and the quadratic approximation of $\log$ Odds Ratios

$$
\begin{equation*}
\mathrm{KI}_{\text {Miettinen }}(\mathrm{OR}):=\left[\mathrm{OR}_{\text {est }} \mathrm{I}^{ \pm \frac{\mathrm{u}_{1-\alpha 2}}{\sqrt{\chi^{2}}}}\right] \tag{2}
\end{equation*}
$$

Under the assumption of Odds Ratios $=1$ or $\approx 1$ the KI of small epidemiological studies can be used for the estimation of Relative Risks.
2. Adjusting age-specific effect as confounder

For generating successive regression models it's indispensable to test the data set on homogeneity. We can suppose diversely likelihoods or Odds for fertility and morbidity status because of different age-specific composition. To avoid a confounding effected overestimation of the truth association it is common to use the Mantel-Haensel-Estimator. Our probands are divided in two different age groups.

$$
\begin{equation*}
\mathrm{OR}_{\text {est }}=\sum_{k=1}^{l} \mathrm{~W}_{\mathrm{k}} \cdot \mathrm{OR}_{\text {est }}, \text { then } \mathrm{W}_{\mathrm{k}}:=\frac{\frac{\mathrm{n}_{10 \mathrm{k}} \cdot \mathrm{n}_{01 \mathrm{k}}}{\mathrm{n} \cdot{ }^{\mathrm{k}}}}{\sum_{k^{\prime}=1}^{l} \frac{\mathrm{n}_{10 \mathrm{k}^{\prime} \cdot \mathrm{n} 01 \mathrm{k}^{\prime}}}{\mathrm{n} \cdots \mathrm{k}^{\prime}}}, \mathrm{k}=1, \ldots, 1 . \tag{3}
\end{equation*}
$$

The stratified analysis is suited because of the adjustments for confounding. The Mantel-HaenselEstimator is a weighted mean of the estimators for every stratum. The weights are the approximation for the reciprocal of the variances of the estimator $\mathrm{OR}_{\text {est. }}$

Under the hypothesis $\mathrm{H}_{0}$ Odds Ratio $=1$ is $\mathrm{OR}_{\text {est_ko }}$ a combined estimator of the stratified results for the common OR following the $X^{2}$ statistics

$$
\begin{align*}
& \text { with } \left.X_{\text {Hom }-1}^{2}=\sum_{k=1}^{l} \frac{\left(\ln \left(\mathrm{OR}_{\text {estk }}\right)-\ln \left(\mathrm{OR}_{\text {est_kombi }}\right)\right)^{2}}{\operatorname{Var}\left(\ln \left(\mathrm{OR}_{\text {estk }}\right)\right)}\right) \\
& \text { then } \operatorname{Var}\left(\ln \left(\mathrm{OR}_{\text {estk }}\right)\right)=\left(\frac{1}{\mathrm{n}_{11 \mathrm{k}}}+\frac{1}{\mathrm{n}_{10 \mathrm{k}}}+\frac{1}{\mathrm{n}_{01 \mathrm{k}}}+\frac{1}{\mathrm{n}_{00 \mathrm{k}}}\right)=\mathrm{vk}^{-1}, k=1, \ldots l . \tag{4}
\end{align*}
$$

By comparing the quadratic differences of log Odds-Ratio-Estimator (observed) in the strata to the combined estimator (expected in case of homogeneity) their distance is purposed to be small in case of homogeneous strata. Following to our context if the distance of the observed Odds for the younger and older cohort and the expected combined Odds containing both subgroups is very small then the strata will be homogeneous.

Then follow $X^{2}$ Hom-1 with hypothesis $H_{0} X^{2}$ - distributed with (l-1) degree of Freedom, the decision (1-a) has to be denied, if

$$
\chi_{\text {Hom-1 }}^{2}>\chi_{\mathrm{t}-1 ; 1-\alpha}^{2}
$$

After the test-statistics on homogeneity are already done we want to focus on the generation of four logistic models. To avoid inconsistent results we use age-standardized variables for fertility and morbidity status because both are affected by events and changes over the life course. Not only the age-specific variation in fertility and morbidity has to be examined, the fertility and morbidity status also might have changed according the cohorts. Hence we are also generating dummies for cohorts to control the cohort-specific variation for our regressors.

$$
\begin{equation*}
\operatorname{logit}(\mathrm{P})=\mathrm{a}+\mathrm{b}_{1} \cdot \mathrm{X}^{(1)}+\mathrm{b}_{2} \cdot \mathrm{X}^{(2)}+\mathrm{b}_{3} \cdot \mathrm{X}^{(3)}, \mathrm{X}^{(1)} \mathrm{X}^{(2)} \mathrm{X}^{(3)}=0,1 \tag{4}
\end{equation*}
$$

P - probality for being infertile/fertile with interval $(0,1)$
$\mathrm{OR}_{1}-\exp \left(\mathrm{b}_{1}\right)$
$\mathrm{OR}_{2}-\exp \left(\mathrm{b}_{2}\right)$
$\mathrm{OR}_{3}-\exp \left(\mathrm{b}_{3}\right)$
$\mathrm{X}^{(1)}=1$, if cohort, cohort1; 0 if cohort, not cohort 1
$X^{(2)}=1$, if cohort, cohort2; 0 if cohort, not cohort 2
$X^{(3)}=1$, if cohort, cohort3; 0 if cohort, not cohort

## 3. Maximum-Likelihood-Estimation

In previous analysis the logistic models couldn't regard definitively the estimations of the probabilities of disease for the total study population. The Maximum-Likelihood-techniques don't consider the individual probabilities, but rather the total probability of disease for all probands.

Referring to the assumption of statistical independency between exposure and non-exposure we are formulating the following. To simplify the maximisation of function $L$ habited by the parameters of the model we use the logarithm of the Log-Likelihood-Function.

$$
\begin{align*}
\mathrm{l}\left(\mathrm{a}, \mathrm{~b}_{1}, \ldots, \mathrm{~b}_{\mathrm{m}}\right) & =\ln \left[\mathrm{L}\left(\mathrm{a}, \mathrm{~b}_{1}, \ldots, \mathrm{~b}_{\mathrm{m}}\right)\right]=\sum_{\mathrm{i}=1}^{\mathrm{n}} \ln \left[\mathrm{P}\left(\mathrm{~K}_{\mathrm{i}}=\mathrm{j} \mid \mathrm{X}_{\mathrm{i}}^{(1)}, \ldots, \mathrm{X}_{\mathrm{i}}^{(\mathrm{m})}\right)\right] \\
& =\sum_{\text {invalid }} \ln \left[1+\exp \left\{-\mathrm{a}-\mathrm{b}_{1} \cdot \mathrm{X}^{(1)} \ldots . \mathrm{b}_{\mathrm{m}} \cdot \mathrm{X}^{(\mathrm{m})}\right\}\right]^{-1}  \tag{5}\\
& =\sum_{\text {healthy }} \ln \left(1-\left[1+\exp \left\{-\mathrm{a}-\mathrm{b}_{1} \cdot \mathrm{X}^{(1)} \ldots . \mathrm{b}_{\mathrm{m}} \cdot \mathrm{X}^{(\mathrm{m})}\right\}\right]^{-1}\right) .
\end{align*}
$$

The Wald-Test is a convenient statistic for simultaneously tests for more than one parameter. This method gives information about the consistence and the real relationship of every regressor and dummy.

$$
\begin{aligned}
& Z_{j}{ }^{2}=\frac{b_{\text {estij }}{ }^{2}}{\operatorname{Var}\left(b_{\text {estij }}\right)}, \mathrm{j}=1, \ldots, \mathrm{~m}, \\
& \text { with } Z^{2}=\left(b_{\text {est1 }}, \ldots, \text { bestm }^{\prime}\right)^{T} \cdot K^{\prime-1} \cdot\left(b_{\text {est1 }}, \ldots, b_{\text {estm }}\right) \text {, } \\
& K^{\prime}=\operatorname{Cov}\left(\text { bestst }, \ldots, \text { bestm' }^{\prime}\right)
\end{aligned}
$$

## Expected Results

With regard to the association between fertility status and lifetime mortality it is expected that differences in lifespan between fertile and infertile men are not the result of a direct causation of infertility on life time mortality in the sense of a biological mechanism. Currently there is no evidence that impaired spermatogenesis might have any direct influence on the lifespan. Differences in mortality between the two groups, i.e. infertile and fertile men can be attributed to differences in morbidity between infertile and fertile men, which are in turn the result of reproductive history, partnership history and health biography. Thus there should be no differences in morbidity and therefore in mortality between subfertile and fertile men under the same conditions.

## Literature

${ }^{i}$ Juul $S$ et al (1999) Regional differences in waiting time to pregnancy: pregancy based surveys from Denmark, France, Germany, Italy and Sweden. The European Infertility and Subfecundity Study Group. Hum Reprod 14 (1250-1254)
${ }^{i i}$ Jorgensen N et al (2002) East-West gradient in semen quality in the Nordic-Baltic area: a study of men from the general population ion Denmark, Norway, Estonia an Finland. Human Reproduction 17,8 (2199-2208)
Carlsen E et al (1992) Evidence for decreasing quality of semen during past 50 Years. BMJ 305 (609613)

Swan SH et al (1997) Have sperm densities declined? A reanalysis of global trend data. Environ Health Perspect 105 (1228-1232)
iii Groos S, Krause W., Mueller UO (2006) Men with subnormal sperm counts live shorter lives. Social Biology, 53, 1-2 (46-60)
Groos S: Lebenszeit-Mortalität von Männern mit normalen und subnormalen
Spermienkonzentrationen. Diss., Marburg 2005
${ }^{\text {iv }}$ Cordier S, Stewart PA: Exposure Assessment. In: Ahrens W, Pigeot I (Eds.): Handbook of Epidemiology. Springer, Berlin - Heidelberg - New York 2005 (437-462)
${ }^{\vee}$ Jacobsen R et al (2000) Risk of testicular cancer in men with abnormal semen characteristics: cohort study. BMJ 321 (789-792)
Coleman MP et al (1993) Trends in cancer incidence and mortality. Lyons: International Agency for Research on Cancer. IARC Scientific Publikation No 121.
Forman D, Moller H (1994) Testicular cancer. Cancer Surv 19-20 (323-341.
Adami HO et al (1994) Testicular cancer in nine northern European countries. Int J Cancer 59 (33-38).

